

**Do Larger Health Insurance Subsidies Benefit Patients or Producers?
Evidence from Medicare Advantage**

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Motivation

- Medicare is the primary source of health insurance for the elderly
 - In 2012, Medicare spending was \$572.5 billion and growing at 4.8%
 - Given the large scale and rapid growth, reforming Medicare is a perpetual policy issue
- One commonly discussed proposal is adjusting subsidies to private Medicare Advantage plans
 - Proponents of larger subsidies argue that increased payments will result in lower premiums / generous benefits
 - Opponents argue that such a move would lead to large profits for insurance companies and health care providers
- At its core, these debates are about economic incidence: Does increasing government subsidies to private Medicare Advantage plans benefit patients or producers?

Background on Medicare

Medicare beneficiaries have two options for hospital + physician coverage:

- Traditional Fee-for-Service Medicare (TM)
 - Public coverage
 - Virtually no provider restrictions
 - Significant patient cost-sharing
- Medicare Advantage (MA)
 - Private coverage
 - Restricted network of providers
 - Little or no patient cost-sharing
 - Often offer supplemental benefits (e.g., vision, dental, drug coverage)

Background on Medicare Advantage

- Medicare eligibles can choose any plan offered in their county
- Plans are given capitation payment from Medicare for each enrolled beneficiary
- Plans can charge a supplemental premium to beneficiaries

Plan payments = capitation payments + premiums

This Paper

- In this paper, we investigate the following questions:
 1. To what degree are increased capitation payments passed through to consumers?
 2. What market factors determine this pass-through rate?

Approach and Findings

- Leverage sharp, differential changes in county-level payments to MA insurers induced by the Benefits Improvement and Protection Act (BIPA) of 2000
- Use this difference-in-differences variation to estimate pass-through
 - For \$1 increase in subsidy, premiums decrease by 45 cents and plan generosity increases by 8 cents
- Write down a simple model to illustrate factors that determine pass-through: selection and market power
- Present empirical evidence on the importance of each of these factors in explaining incomplete pass-through

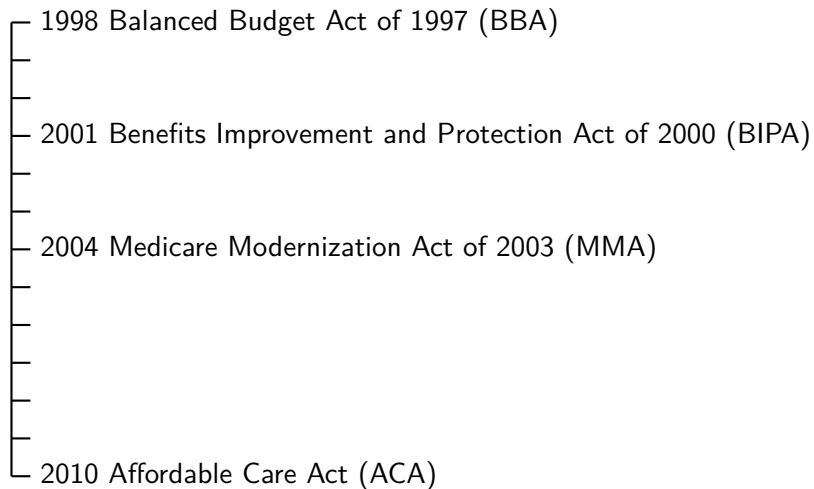
Related Literature

- Pass-through in MA
 - Duggan, Starc, and Vabson (2015)
 - Song, Landrum, and Chernew (2013)
- Selection into MA
 - "Switcher" studies (e.g., Brown et al. 2014; Newhouse et al. 2012)
 - We use exogenous variation in prices (e.g. Einav, Finkelstein, Cullen 2010)
- Market power in health insurance
 - Curto et al. (2015) on market power in MA
 - Dafny (2010) and Dafny et al. (2012) on market power in ESI

Outline

- **Background**
- Research design
- Pass-through
- Model
- Selection and market power

An Abridged History of Payment Reforms



MA Payments

Capitation payments intended to reflect counterfactual TM costs

$$\text{Capitation payment}_{ijt} = b_{jt} \times r_{it}$$

- b_{jt} is county-level “base payment”
 - Pre BIPA, largely determined by historical average TM costs
 - Base payments increased by approx 2% per year
- r_{it} is demographic risk adjustment
 - Normalized to have mean 1 in entire population
 - Comprehensive risk adjustment introduced in 2004

Data

- Multiple sources:
 - MA Rate-books: Payments for county \times year
 - Plan Service Files: Benefits and premiums by plan \times year
 - CMS Beneficiary Summary File: admin cost data for TM
 - CMS Denominator File: admin demographic data for all Medicare
- Time frame: 1997-2003
 - Premium data for 1997-2003
 - Benefits data for 2000-2003
 - Plan quality data for 1999-2003
 - Costs data for 1999-2003

Sample Construction

- Aggregate data to county \times year panel
 - Weight plan-level attributes by enrollment shares
 - Weight county \times years by number of beneficiaries in each county
- Only observe plan attributes when 1+ plan in county
 - Baseline: County \times years with 1+ plan
 - Show that variation does not affect entry / exit into sample

Summary Statistics

Table: All Counties, 1997-2003

	Mean	Std. Dev.	Min.	Max.
Base Payment (\$ per month)	490.58	83.96	222.99	777.91
At Least One Plan	65.1%	47.7%	0%	100%
Number of Plans	1.78	1.73	0	7
MA Enrollment	19.1%	18.4%	0%	69.8%
TM Costs (\$ per month)	486.53	103.94	136.87	940.08

Summary Statistics

Table: County \times Years with At Least One Plan, 1997-2003

	Mean	Std. Dev.	Min.	Max.
County-Level Premium (\$ per month)				
Mean	22.71	27.82	0	156.29
Min	15.05	26.25	0	156.29
Median	21.60	29.60	0	156.29
Max	33.56	33.54	0	194.47
County-Level Benefits*				
Physician Copay (\$ per visit)	7.89	4.95	0	56.15
Specialist Copay (\$ per visit)	14.39	6.79	0	95.72
Drug Coverage	70.5%	41.1%	0%	100%
Dental Coverage	27.4%	35.7%	0%	100%
Vision Coverage	69.9%	39.8%	0%	100%
Hearing Aid Coverage	40.0%	42.1%	0%	100%
Number of Plans	2.75	1.41	1	7
HHI	5,696	2,584	1,778	10,000
MA Enrollment	28.8%	16.1%	1.1%	67.6%
TM Costs (\$ per month)	521.80	106.65	254.96	940.08

*Benefits data are only available for 2000-2003

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MA Payments and BIPA

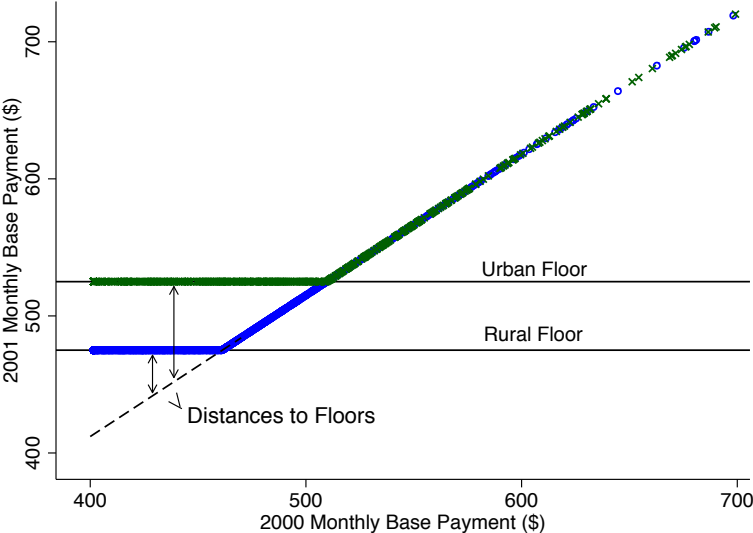
- Benefits Improvement and Protection Act of 2000
 - Implemented rural and urban payment floors*
- Base payments

$$b_{jt} = \begin{cases} \tilde{c}_{jt} & \text{if } t < 2001 \\ \max \{ \tilde{c}_{jt}, \underline{b}_{u(j)t} \} & \text{if } t \geq 2001, \end{cases}$$

- \tilde{c}_{jt} is the base payment absent the BIPA floors
- $\underline{b}_{u(j)t}$ is the relevant urban or rural payment floor

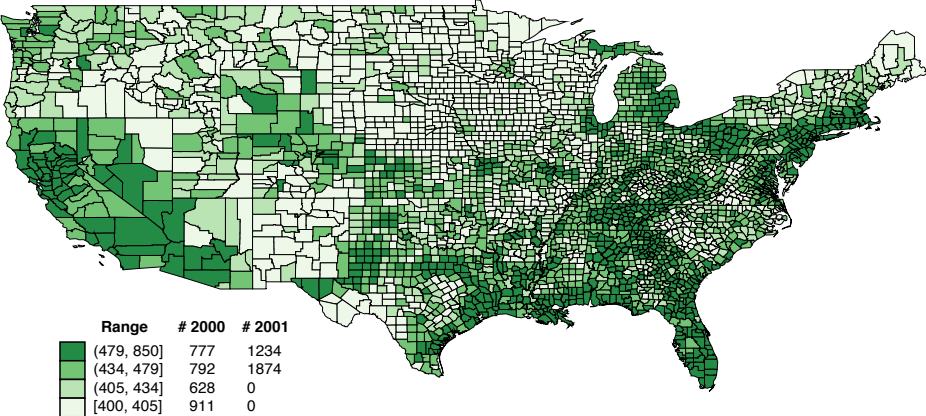
*Required plans to submit new premiums and benefits to take effect in February 2001. We define 2001 premiums using these post-update value

BIPA Payment Floors



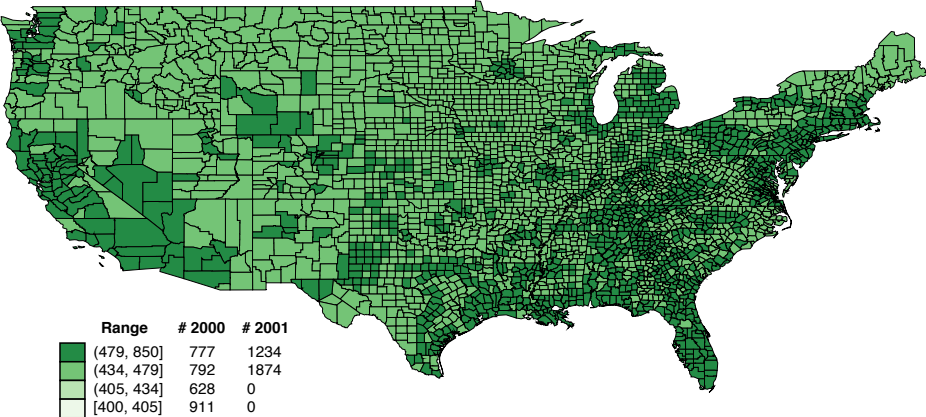
Effect of BIPA on Payments

Figure: Pre-BIPA Payments, 2000



Effect of BIPA on Payments

Figure: Post-BIPA Payments, 2001



Payment Floors

	Mean	Std. Dev.	Percentiles		
			25th	50th	75th
Non-Floor County (N = 886)					
Δ Base Payment	14.39	1.58	13.17	14.03	15.10
% Change in Base Payment	3.0%	0.0%	3.0%	3.0%	3.0%
Rural Floor County (N = 1,831)					
Δ Base Payment	52.94	17.16	39.67	62.59	67.18
% Change in Base Payment	14.1%	4.9%	10.0%	16.8%	18.3%
Urban Floor County (N = 426)					
Δ Base Payment	64.67	29.56	38.90	62.33	89.05
% Change in Base Payment	16.1%	8.4%	8.8%	14.9%	22.7%

Econometric Model

- Measure exposure to BIPA with a *distance-to-floor* measure

$$\Delta b_{jt} = \max \left\{ \tilde{b}_{u(j)t} - \tilde{c}_{jt}, 0 \right\}$$

- $\tilde{b}_{u(j)t}$ is relevant urban/rural floor in year t
- \tilde{c}_{jt} is payment rate in absence of the floor in county j in year t

► More Details

Econometric Model

- Difference-in-differences with year-specific coefficients

$$y_{jt} = \alpha_j + \alpha_t + \left(\sum_{t \neq 2000} \beta_t \times I_t \times \Delta b_{jt} \right) + f(X_{jt}) + \epsilon_{jt}$$

- α_j and α_t are county and year fixed effects
- $f(X_{jt})$ is a flexible set of controls
- Normalize $\beta_{2000} = 0$ in year when BIPA was passed
- Cluster standard errors at the county level

Identification

Assumption: In the absence of BIPA, outcomes for counties that were differentially affected by the payment floors would have evolved in parallel

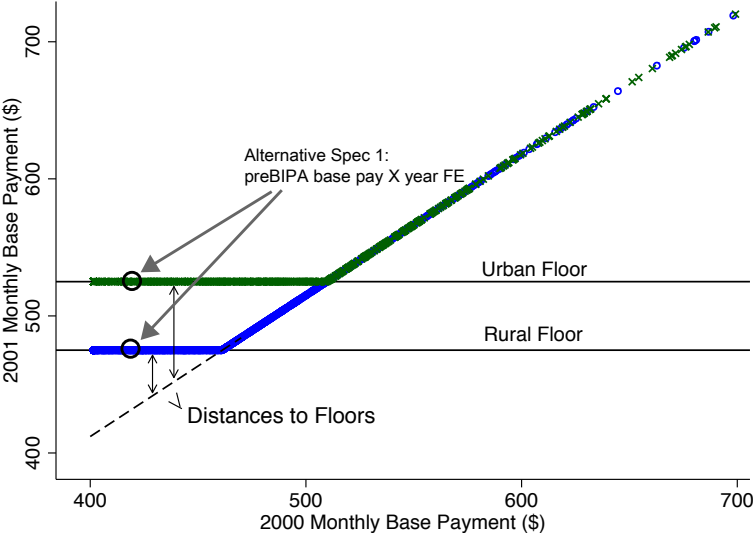
- Two broad approaches to assessing the validity of this assumption:
 - Plot β_t 's over time to visually inspect for spurious pre-existing trends
 - Show results robust to alternative specifications that isolate two complementary sources of identifying variation
 1. Include pre-BIPA Base Payment X Year FE
 2. Include Urban X Year FE

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BIPA Payment Floors

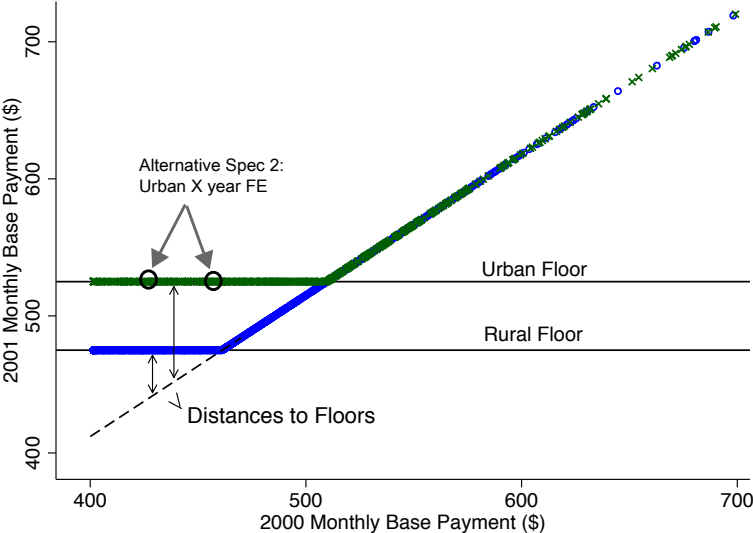


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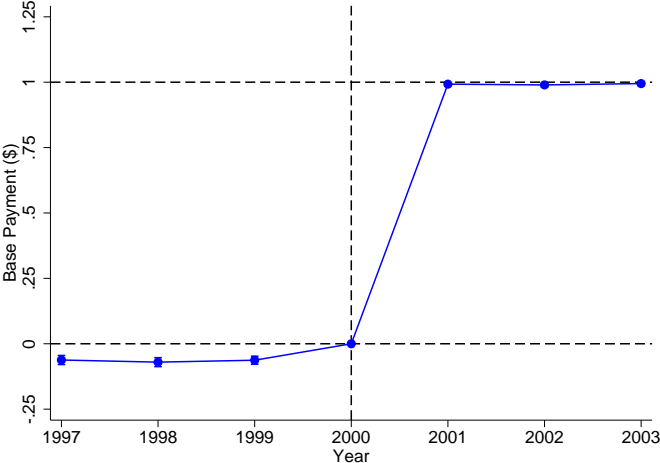
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BIPA Payment Floors



First Stage Impact on Base Payment

Figure: Impact of \$1 Increase in Distance to Floor



First Stage, Alternative Specifications

Figure: Impact of \$1 Increase in Distance to Floor

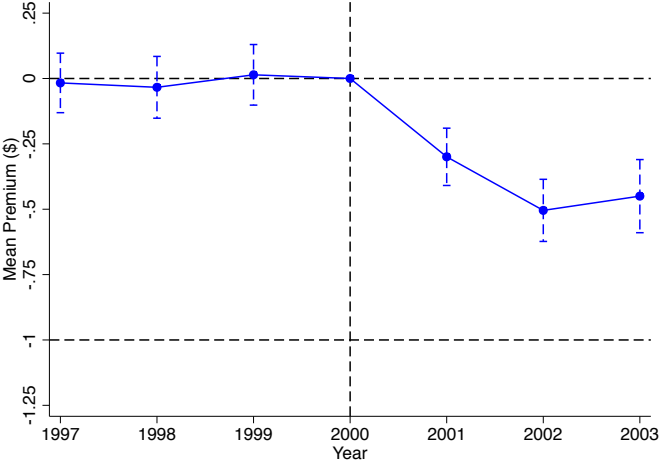
	Dependent Variable: Base Payment (\$)		
	(1)	(2)	(3)
Δb X 2001	0.993 (0.003)	0.996 (0.004)	0.993 (0.003)
Δb X 2002	0.990 (0.004)	0.997 (0.005)	0.987 (0.004)
Δb X 2003	0.995 (0.004)	1.002 (0.005)	0.992 (0.004)
Main Effects			
County FE	X	X	X
Year FE	X	X	X
Additional Controls			
Pre-BIPA Payment X Year FE		X	
Urban X Year FE			X
Pre-BIPA Mean of Dep. Var.	515.15	515.15	515.15
R-Squared	1.000	1.000	1.000

Outline

- Background and data
- Research design
- **Pass-through**
- Model
- Selection and market power

Mean Premiums

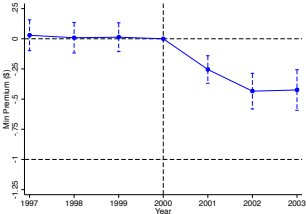
Figure: Impact of \$1 Increase in Monthly Payments



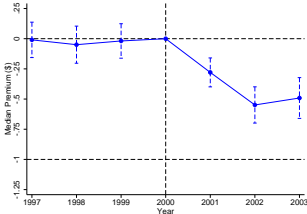
Distribution of Premiums

Figure: Impact of \$1 Increase in Monthly Payments

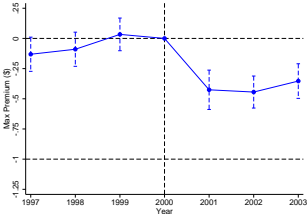
(a) Min



(b) Median



(c) Max



Premiums Robustness

For every \$1 increase in subsidy, mean premiums decline by 45 cents

Obtain similar estimates when...

1. Investigate effect on distribution of premiums
2. Estimate alternative specifications that isolate subsets of identifying variation [▶ Subsets of variation](#)
3. Estimate Tobit specifications that take into account that plans could not give rebates during our time period [▶ Tobit regressions](#)
4. Aggregate up to a higher level [▶ Aggregated regressions](#)
5. Examine detailed timing using monthly data [▶ Monthly regressions](#)

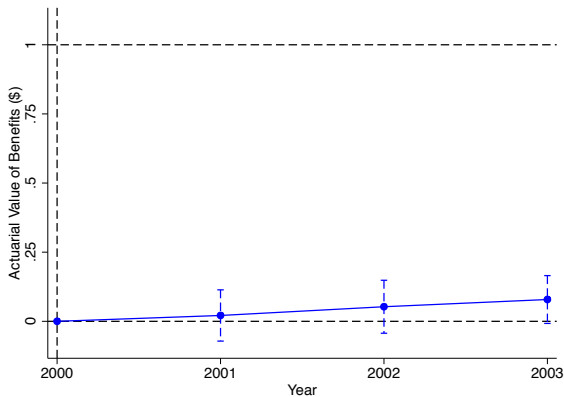
Benefits

Insurers could have alternatively passed-through subsidies via benefits

- We evaluate the impact on benefits using multiple approaches:
 1. Impact of \$50 increase ($\sim 10\%$) in payments on copays, dental, etc.
[▶ Additional Figures](#)
 2. Impact on actuarial value using data on utilization / insurance payments from MEPS

Monetized Benefits

Figure: Impact of \$1 Increase in Monthly Payments



- By 2003, max pass-through in benefits of 8 cents on the dollar

Unobserved Quality

Limited concern in this setting for two reasons

1. Rich product characteristics data

- We see everything consumers see at the point of sale
- Many other characteristics significantly constrained by regulation (e.g., essential benefits, network adequacy)

2. Additional analysis of quality data [▶ Quality Analysis](#)

- Precisely estimated zero on beneficiary's subjective evaluations of plan quality (CAHPS)
- Precisely estimated zero on clinical quality measures (HEDIS)

Plan Availability

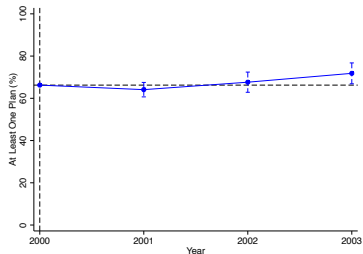
Examine two margins

- Extensive: Percent of counties with at least one plan
- Intensive: HHI conditional on having at least one plan

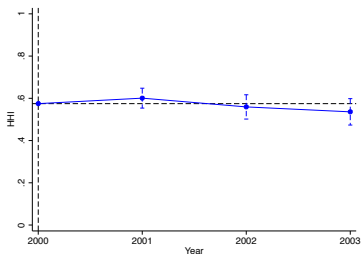
Plan Availability: Extensive and Intensive Margins

Figure: Impact of \$50 Increase in Monthly Payments

(a) Extensive Margin (1+ plan)



(b) Intensive Margin, (HHI)



► Plan Availability Table

Pass-through Estimates: Key Takeaways

For every \$1 marginal increase in subsidy:

- 45 cents passed-through in lower premiums
 - 8 cents passed-through in more generous benefits
 - No detectable effect on entry
- ⇒ About one-half (53 cents) of increase flows to consumers, with 95% confidence interval (35 cents, 71 cents)

Outline

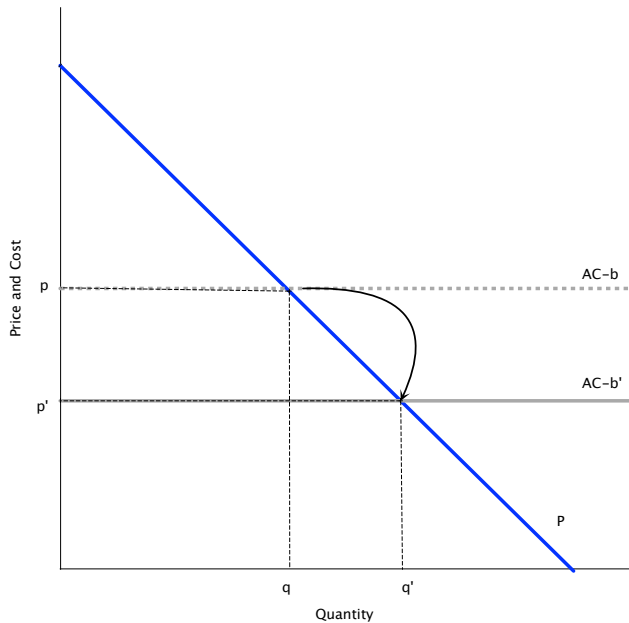
- Background and data
- Research design
- Pass-through
- **Model**
- Selection and market power

Approach

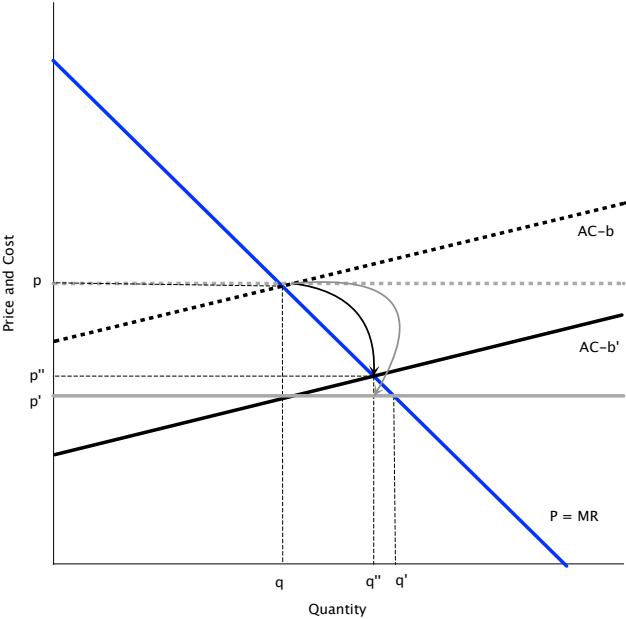
⇒ Potential Mechanisms: Advantageous Selection and Market Power

- Graphical intuition
- Model that relates pass-through to these factors

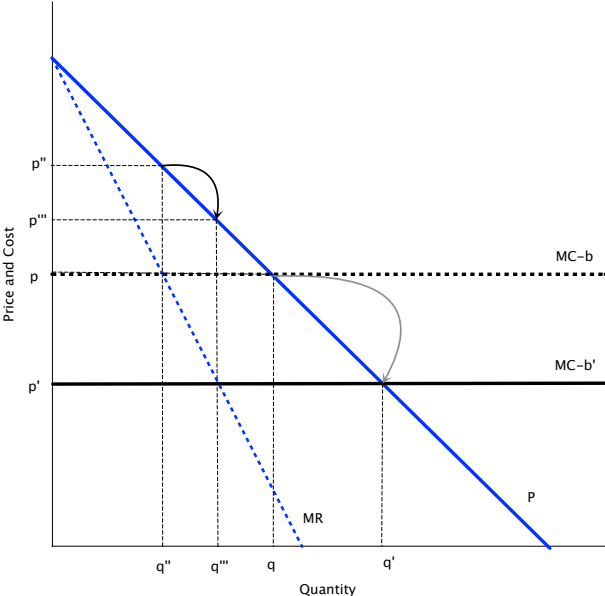
No Selection, Perfect Competition



Advantageous Selection, Perfect Competition



No Selection, Monopoly



Model Setup

Build more general model that expresses pass-through as a function market power and selection

- Aggregate demand: $Q(p) \in [0, 1]$
- Aggregate costs for industry: $C(Q) \equiv \int_{v_i \geq p^{-1}(Q)} c_i$
 - Average costs: $AC(Q) \equiv \frac{C(Q)}{Q}$
 - Marginal costs: $MC(Q) \equiv C'(Q)$
- Selection
 - Adverse selection: $MC'(Q) < 0$
 - Advantageous selection: $MC'(Q) > 0$

Equilibrium

- Perfect competition characterized by zero profits

$$p = AC(Q) - b$$

- Monopolist's first order condition

$$p = \mu(p) + MC(Q) - b$$

- $\mu(p) \equiv -\frac{Q(p)}{Q'(p)}$ is absolute markup term

Market Power

Following Weyl-Fabinger (2013), introduce conduct parameter $\theta \in [0, 1]$

$$p = \theta(\mu(p) + MC(Q) - b) + (1 - \theta)(AC(Q) - b)$$

- Nests extremes
 - Perfect competition: $\theta = 0$. Monopoly: $\theta = 1$
- Reduced form of standard models
 - Cournot: $\theta = 1/n$
 - Diff product Bertrand: $\theta = 1 - \text{aggregate diversion ratio}$
 - Requires “symmetry assumptions” on selection (see Mahoney and Weyl, 2014)

Pass-Through

- Define pass-through as $\rho \equiv -\frac{dp}{db}$
- Fully differentiating FOC yields

$$\rho = \frac{1}{1 - (1 - \theta) \left(\frac{dAC}{dp} \right) - \theta \left(\frac{d\mu}{dp} + \frac{dMC}{dp} \right)}$$

- Assuming linear demand and costs

$$\rho = \underbrace{\left(\frac{1}{1 - \frac{dAC}{dp}} \right)}_{\text{Selection}} \underbrace{\left(\frac{1}{1 + \theta} \right)}_{\text{Market power}}$$

Outline

- Background and data
- Research design
- Pass-through
- Model
- **Selection and market power**

Impact of Selection

- Want to estimate

$$\tilde{\rho} = \frac{1}{1 - \frac{dAC}{dp}}$$

- Two interpretations
 1. Reduction in pass-through due to selection in perfect comp baseline
 2. Proportional reduction in pass-through in linear model with any level of competition

Impact of Selection

Introducing risk rating

$$\tilde{p} = \frac{AR}{1 - \left(\frac{dAC}{dp} - b \frac{dAR}{dp} \right)}$$

- $\frac{dAC}{dp} - b \frac{dAR}{dp}$ measures selection *net of risk adjustment payments*
- Scaled by AR to convert base payment into capitation payment

Estimation Approach

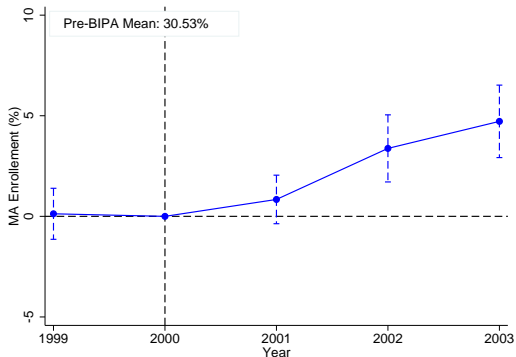
- Main challenge: Have admin data on TM costs, not MA plan costs
 - Prior literature looks at switchers: Do beneficiaries who switch from FFS to MA have lower $t - 1$ costs than beneficiaries who stay?
 - Evidence is mixed (e.g., Brown et al. 2014; Newhouse et al. 2012)
 - Magnitudes are not economically interpretable
 - Does not identify selection with respect to premiums

Estimation Approach

- Our approach builds on / formalizes switcher idea with two assumptions:
 - A1.** Costs under MA and TM are proportional $c_i^{MA}/c_i^{TM} = \phi$ with $\phi \leq 1$
 - $\phi \leq 1$ consistent with Bundorf Levin Mahoney (2012), other evidence on managed care vs. fee for service cost structures
 - A2.** Cost curves are linear so that selection is parameterized by single slope parameter
- Under these assumptions
 - TM slope provides upper bound on MA slope and therefore explanatory power of selection

MA Enrollment

Figure: Impact of \$50 Increase in Monthly Payment

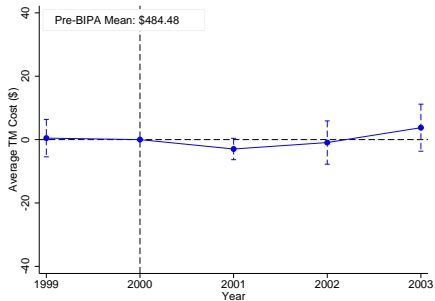


- \$23 decrease in premiums raises MA by 4.7 pp on base of 30.5%

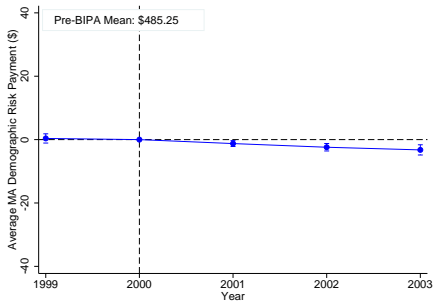
Average Costs

Figure: Impact of \$50 Increase in Monthly Payment

(a) Average TM Costs



(b) Risk Adjustment Payments



- Slope of $\frac{dAC^{MA}}{dQ} - b \frac{dAR^{MA}}{dQ}$ is \$149 with 95% CI of (-\$9, \$307)
- No effect on utilization [▶ Evidence on Utilization](#)

Impact of Market Power

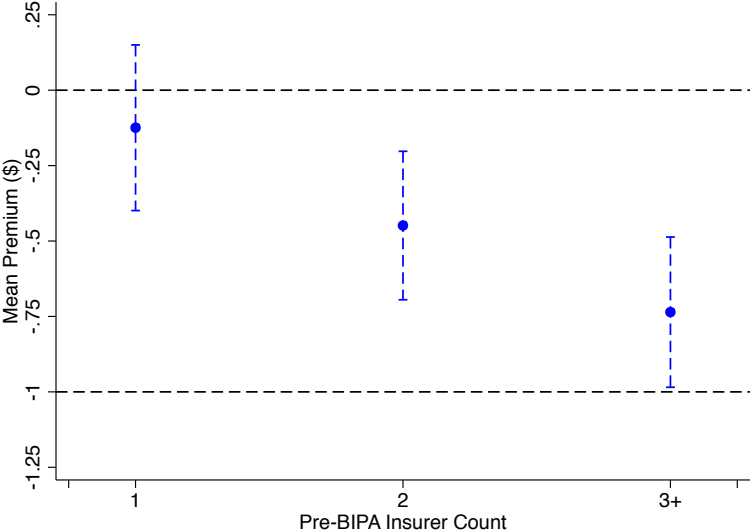
- Estimates above imply that $\tilde{\rho} = 85$ cents

▶ Table of Estimates

- Theory: Residual ≈ 35 ppt due to market power
- **Can we find supporting empirical evidence?**
- Idea: Heterogeneity in pass-through estimates by pre-BIPA measures of market power
 - Number of pre-BIPA insurance plans
 - Pre-BIPA Insurer HHI

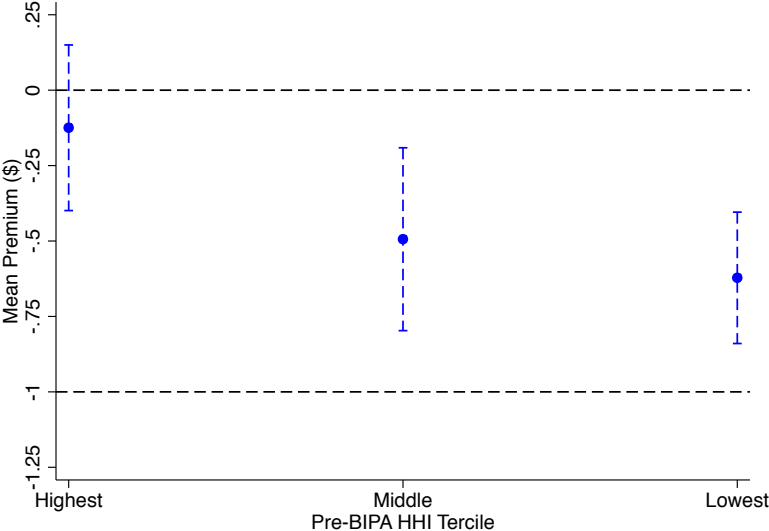
Heterogeneity by pre-BIPA Number of Insurers

Figure: Pass-through



Heterogeneity by pre-BIPA Insurer HHI

Figure: Pass-through



Conclusion

- Used sharp, differential increase in MA payments to study allocation of (marginal) surplus in privatized Medicare
 - One-half of increase passed-through to consumers
 - ⇒ Implications for \$156B in MA payment reductions scheduled under ACA
- Investigate explanations of incomplete pass-through
 - Advantageous selection has limited explanatory power
 - Evidence suggests market power more likely explanatory factor
 - ⇒ Implication is that efforts to make markets more competitive may be key to increasing consumer surplus on the margin

- Measure exposure to BIPA with *distance-to-floor variable*:

$$\Delta b_{jt} = \max \left\{ \tilde{b}_{u(j)t} - \tilde{c}_{jt}, \quad 0 \right\},$$

- Use data on base rates in the pre-period to construct \tilde{c}_{jt} , the monthly payment in the absence of the floor

$$\tilde{c}_{jt} = \begin{cases} c_{jt} & \text{if } t \leq 2001 \\ c_{j,2001} \cdot 1.02^{(t-2001)} & \text{if } t > 2001 \end{cases}$$

- Use data on floors in the post-period to construct \tilde{b}_{jt} , the counterfactual urban or rural payment floors:

$$\tilde{b}_{u(j)t} = \begin{cases} \underline{b}_{u(j),2001} \cdot 1.02^{(t-2001)} & \text{if } t < 2001 \\ \underline{b}_{u(j)t} & \text{if } t \geq 2001 \end{cases}$$

Premiums, Alternative Specifications

Table: Impact of \$1 Increase in Monthly Payments

	Dependent Variable:		
	Mean Monthly Premium (\$)		
	(1)	(2)	(3)
Δb X 2001	-0.301 (0.056)	-0.178 (0.095)	-0.314 (0.057)
Δb X 2002	-0.503 (0.061)	-0.352 (0.112)	-0.516 (0.061)
Δb X 2003	-0.444 (0.072)	-0.378 (0.120)	-0.445 (0.073)
Main Effects			
County FE	X	X	X
Year FE	X	X	X
Additional Controls			
Pre-BIPA Payment X Year FE		X	
Urban X Year FE			X
Pre-BIPA Mean of Dep. Var.	12.10	12.10	12.10
R-Squared	0.71	0.71	0.71

Premium Regressions, Plan Level Regressions

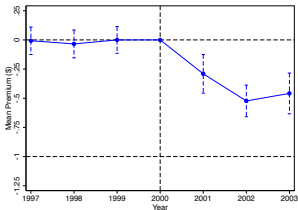
Table: Impact of \$1 Increase in Monthly Payments

	Dependent Variable: Monthly Premium (\$)					
	Linear Regression			Tobit Regression		
	(1)	(2)	(3)	(4)	(5)	(6)
$\Delta b \times 2001$	-0.298 (0.056)	-0.195 (0.094)	-0.311 (0.056)	-0.461 (0.011)	-0.181 (0.016)	-0.485 (0.011)
$\Delta b \times 2002$	-0.502 (0.060)	-0.440 (0.112)	-0.514 (0.060)	-0.577 (0.008)	-0.370 (0.011)	-0.586 (0.008)
$\Delta b \times 2003$	-0.447 (0.071)	-0.424 (0.123)	-0.449 (0.072)	-0.537 (0.010)	-0.380 (0.012)	-0.539 (0.010)
Main Effects						
County FE	X	X	X	X	X	X
Year FE	X	X	X	X	X	X
Additional Controls						
Pre-BIPA Payment X Year FE		X			X	
Urban X Year FE			X			X
Pre-BIPA Mean of Dep. Var.	12.56	12.56	12.56	12.56	12.56	12.56
R-Squared	0.60	0.60	0.60	N/A	N/A	N/A

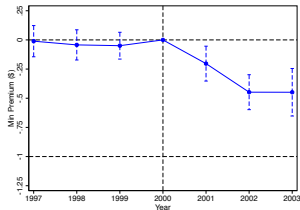
Unit of observation aggregated to MSA \times state \times year

Figure: Impact of \$1 Increase in Monthly Payments

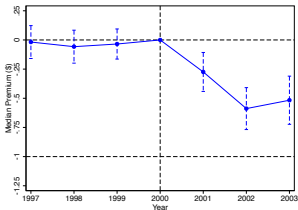
(a) Mean



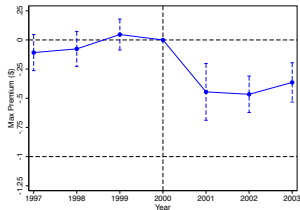
(b) Min



(c) Median

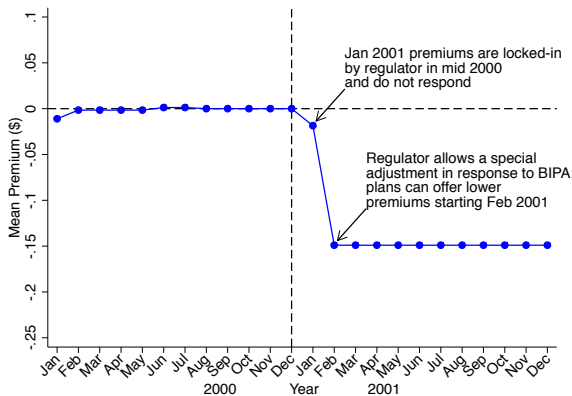


(d) Max



Detailed Timing of Effects

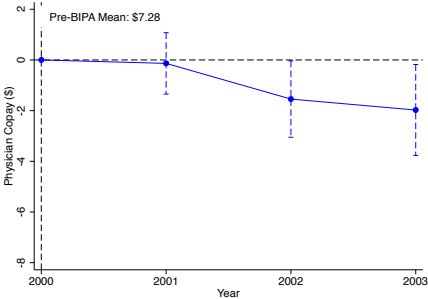
Figure: Impact of \$1 Increase in Monthly Payments



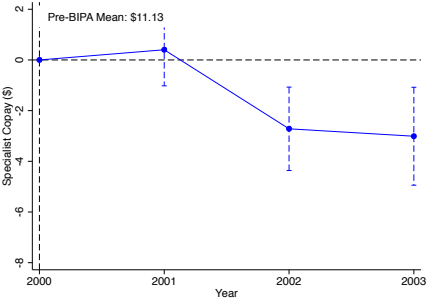
Benefits: Average Copays

Figure: Impact of \$50 Increase in Monthly Payments

(a) Physician



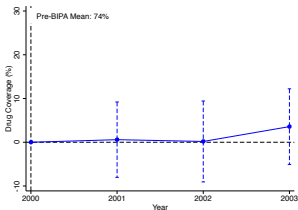
(b) Specialist



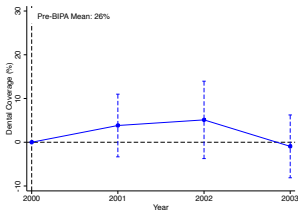
Benefits: Drugs, Dental, Vision, Hearing Aid Coverage

Figure: Impact of \$50 Increase in Monthly Payments

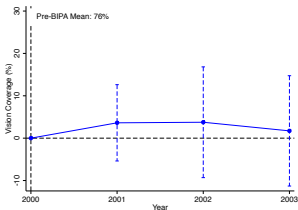
(a) Drugs



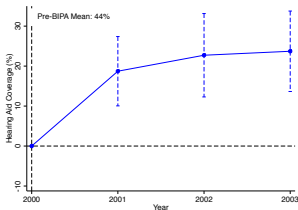
(b) Dental



(c) Vision



(d) Hearing Aid



Benefits Regressions

Table: Impact of Increase in Monthly Payments

	Dependent Variable:						
	Physician Copay (\$)	Specialist Copay (\$)	Drug Coverage (%)	Dental Coverage (%)	Vision Coverage (%)	Hearing Aid Coverage (%)	Actuarial Value (\$)
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Δb X 2001*	-0.136 (0.618)	0.402 (0.726)	0.589 (4.396)	3.827 (3.654)	3.622 (4.595)	18.725 (4.424)	0.021 (0.047)
Δb X 2002*	-1.544 (0.769)	-2.717 (0.840)	0.180 (4.719)	5.111 (4.513)	3.756 (6.668)	22.721 (5.321)	0.053 (0.049)
Δb X 2003*	-1.976 (0.917)	-3.010 (0.986)	3.571 (4.410)	-0.939 (3.664)	1.721 (6.643)	23.712 (5.132)	0.079 (0.044)
Main Effects							
County FE	X	X	X	X	X	X	X
Year FE	X	X	X	X	X	X	X
Pre-BIPA Mean of Dep. Var.	7.28	11.13	74.20	26.11	75.84	44.44	n/a
R-Squared	0.66	0.70	0.83	0.68	0.75	0.85	0.83

*Final column displays the effect of a \$1 increase in monthly payments. All other columns display the impact of a \$50 increase in monthly payments.

[▶ Back to Monetized Benefits](#)

Benefits Regressions, Additional Specifications

Table: Impact of \$50 Increase in Monthly Payments

	Dependent Variable:													
	Physician Copay (\$)		Specialist Copay (\$)		Drug Coverage (%)		Dental Coverage (%)		Vision Coverage (%)		Hearing Aid Coverage (%)		Actuarial Value (\$)	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)
Δb X 2001*	-0.24 (0.67)	-0.12 (0.63)	0.44 (0.83)	0.46 (0.73)	4.45 (4.73)	0.94 (4.41)	7.84 (5.07)	4.19 (3.77)	3.82 (5.80)	3.77 (4.68)	18.99 (5.35)	18.66 (4.51)	0.07 (0.05)	0.02 (0.05)
Δb X 2002*	-1.69 (0.84)	-1.70 (0.78)	-2.88 (1.01)	-2.78 (0.85)	4.47 (5.15)	0.72 (4.83)	12.41 (5.62)	6.62 (4.58)	8.06 (7.30)	3.85 (6.71)	26.13 (6.34)	22.74 (5.46)	0.11 (0.06)	0.06 (0.05)
Δb X 2003*	-2.78 (1.01)	-2.14 (0.93)	-3.10 (1.27)	-3.21 (1.01)	3.86 (4.77)	4.92 (4.48)	-0.62 (5.11)	0.73 (3.66)	6.10 (7.34)	1.77 (6.69)	21.86 (6.55)	23.79 (5.26)	0.09 (0.05)	0.10 (0.04)
Main Effects														
County FE	X	X	X	X	X	X	X	X	X	X	X	X	X	X
Year FE	X	X	X	X	X	X	X	X	X	X	X	X	X	X
Additional Controls														
Pre-BIPA Base Payment X Year FE	X		X		X		X		X		X		X	
Urban X Year FE		X		X		X		X		X		X		X
Pre-BIPA Mean of Dep. Var.	7.28	7.28	11.13	11.13	74.20	74.20	26.11	26.11	75.84	75.84	44.44	44.44	35.95	35.95
R-Squared	0.67	0.66	0.70	0.70	0.83	0.83	0.69	0.68	0.76	0.75	0.85	0.85	0.83	0.83

*Final column displays the effect of a \$1 increase in monthly payments. All other columns display the impact of a \$50 increase in monthly payments. [▶ Back to Monetized Benefits](#)

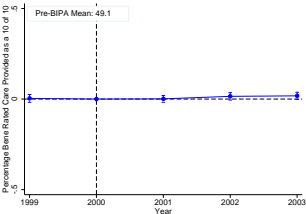
Plan Quality

- Measures of plan quality (Dafny and Dranove, 2008)
 1. Measures listed in *Medicare & You* booklet
 - Quality of care, quality of doctor communication from CAHPS, mammogram rate from HEDIS
 2. Unreported quality index
 - Beta blockers, diabetic eye exams, preventive routine exams from HEDIS

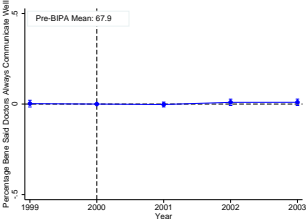
Plan Quality

Figure: Impact of \$50 Increase in Payment Floor

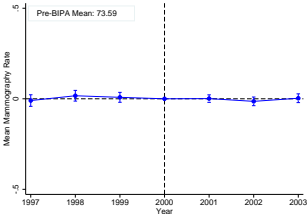
(a) Quality of Care



(b) Doctor Communication

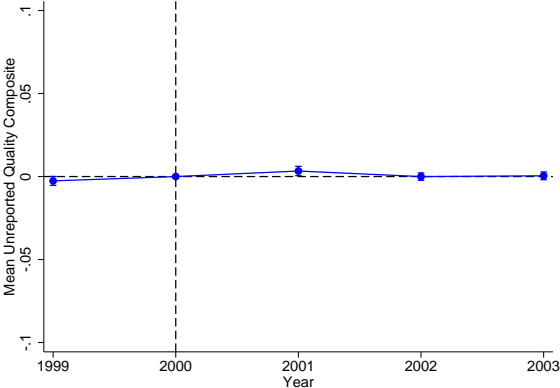


(c) Mammography



Unreported Quality Index

Figure: Impact of \$50 Increase in Monthly Payments



Standardized composite of beta blockers, preventive care visits, diabetic eye exams

Plan Availability, Alternative Specifications

Table: Impact of \$50 Increase in Monthly Payments

	Dependent Variable:					
	At Least One Plan (%)			HHI*		
	(1)	(2)	(3)	(4)	(5)	(6)
Δb X 2001	-2.15 (1.75)	0.04 (2.06)	-2.34 (1.76)	0.037 (0.030)	-0.031 (0.033)	0.039 (0.030)
Δb X 2002	1.39 (2.44)	2.92 (2.65)	1.92 (2.46)	-0.001 (0.034)	-0.056 (0.037)	-0.012 (0.035)
Δb X 2003	5.58 (2.52)	7.89 (2.91)	6.11 (2.55)	-0.030 (0.037)	-0.097 (0.040)	-0.043 (0.038)
Main Effects						
County FE	X	X	X	X	X	X
Year FE	X	X	X	X	X	X
Additional Controls						
Pre-BIPA Base Payment X Year FE		X			X	
Urban X Year FE			X			X
Pre-BIPA Mean of Dep. Var.	66.2	66.2	66.2	0.51	0.51	0.51
R-Squared	0.91	0.91	0.91	0.77	0.78	0.77

Estimation Approach Details

- Proportional costs imply proportional costs for marginal individual

$$MC^{MA}(Q^{MA}) = \phi MC^{TM}(Q^{TM})$$

- Because $Q^{TM} = 1 - Q^{MA}$, slopes under MA and TM are of reversed sign and proportional

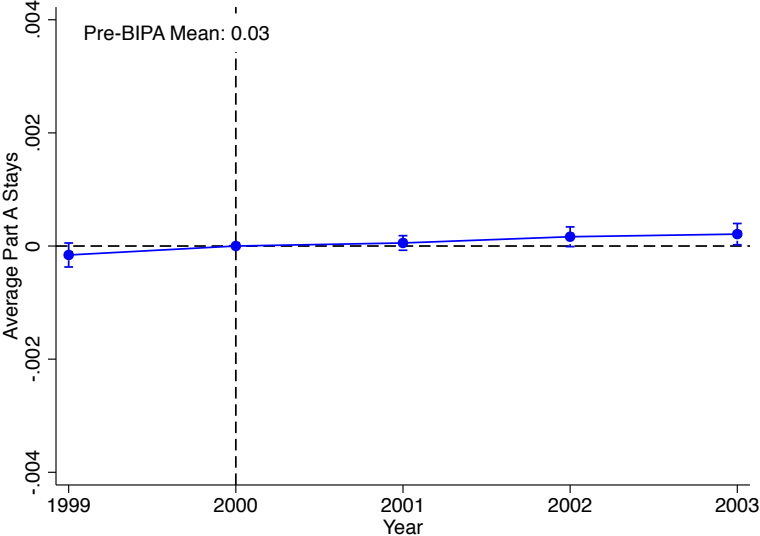
$$\frac{dMC^{MA}}{dQ^{MA}} = -\phi \frac{dMC^{TM}}{dQ^{TM}}$$

- Applying linearity to translate from MC to AC yields

$$\frac{dAC^{MA}}{dQ^{MA}} = -\phi \frac{dAC^{TM}}{dQ^{TM}}$$

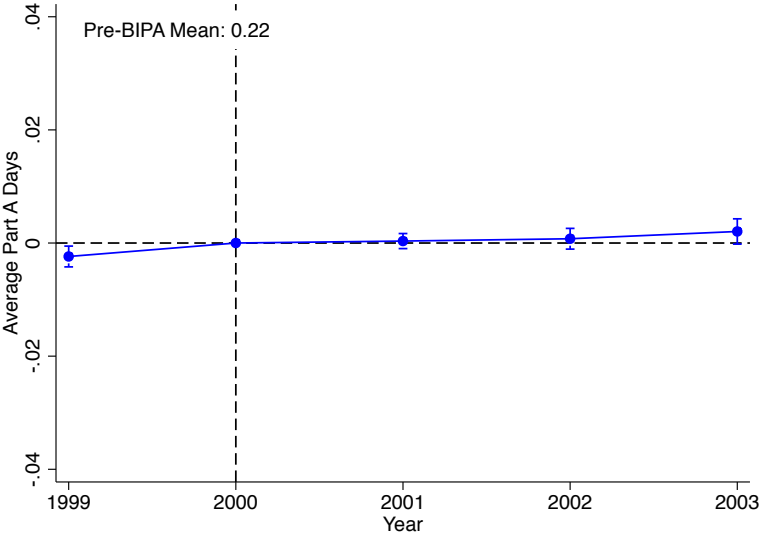
Part A Stays

Figure: Impact of \$50 Increase in Monthly Payments



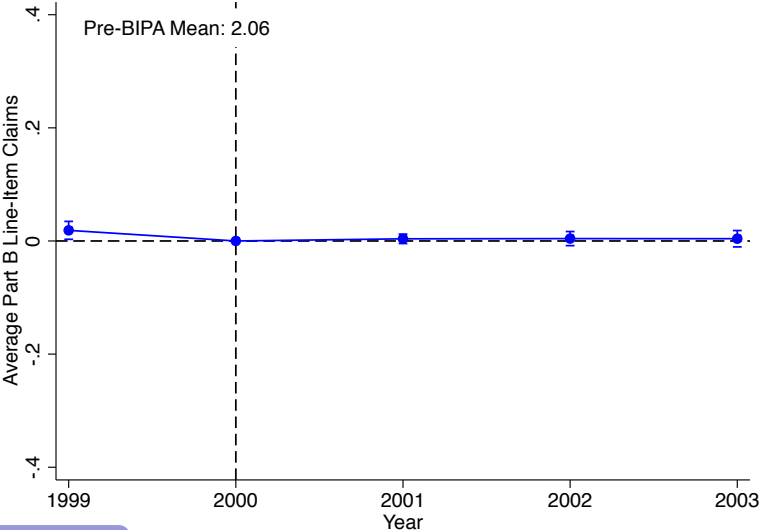
Part A Days

Figure: Impact of \$50 Increase in Monthly Payments



Part B Line-Item Claims

Figure: Impact of \$50 Increase in Monthly Payments



Selection Regression Estimates

Table: Impact of \$50 Increase in Monthly Payment

	Dependent Variable:				Implied Pass-Through with Selection (p)
			MA Risk Adjustment	Mean Premiums*	
	MA Enrollment (%)	TM Costs (\$)	(\$)	(\$)	
	(1)	(2)	(3)	(4)	(5)
Panel A: Yearly BIPA Effect					
Δb X 2001	0.84 (0.62)	-2.96 (1.72)	-1.25 (0.47)	-0.300 (0.056)	1.076 (0.267)
Δb X 2002	3.38 (0.85)	-0.93 (3.48)	-2.41 (0.60)	-0.504 (0.061)	0.903 (0.125)
Δb X 2003	4.72 (0.92)	3.76 (3.79)	-3.24 (0.82)	-0.450 (0.071)	0.732 (0.103)
Panel B: Pooled Post-BIPA Effect					
Δb X Post-BIPA	3.27 (0.73)	0.21 (2.86)	-2.68 (0.60)	-0.44 (0.05)	0.845 (0.095)
Controls: All Panels					
Main Effects					
County FE	X	X	X	X	
Year FE	X	X	X	X	
Pre-BIPA Mean of Dep. Var.	30.53	485.25	484.48	10.90	

*Column 4 displays the impact of a \$1 increase in monthly payments; all other columns display the effect of a \$50 increase in monthly payments.

[▶ Additional Specifications](#)

[▶ Back to Selection Section](#)

Selection Regression Estimates, Additional Specifications

Table: Impact of \$50 Increase in Monthly Payments

	Dependent Variable:								
	MA Enrollment (%)			TM Costs (\$)			MA Risk Adjustment (\$)		
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Panel A: Yearly BIPA Effect									
Δb X 2001	0.84 (0.62)	2.26 (0.68)	0.83 (0.63)	-2.96 (1.72)	3.04 (1.94)	-3.22 (1.78)	-1.25 (0.47)	-0.75 (0.91)	-1.35 (0.50)
Δb X 2002	3.38 (0.85)	5.17 (0.96)	3.65 (0.86)	-0.93 (3.48)	5.34 (3.96)	-1.19 (3.59)	-2.41 (0.60)	-2.76 (1.09)	-2.50 (0.61)
Δb X 2003	4.72 (0.92)	7.31 (1.04)	5.08 (0.93)	3.76 (3.79)	10.84 (5.25)	3.74 (3.91)	-3.24 (0.82)	-3.25 (1.28)	-3.36 (0.84)
Panel B: Pooled Post-BIPA Effect									
Δb X Post-BIPA	3.27 (0.73)	5.95 (0.86)	3.47 (0.74)	0.21 (2.86)	8.18 (3.53)	0.15 (2.98)	-2.68 (0.60)	-2.47 (1.06)	-2.80 (0.62)
Panel C: Pooled Post-BIPA Effect									
Main Effects									
County FE	X	X	X	X	X	X	X	X	X
Year FE	X	X	X	X	X	X	X	X	X
Additional Controls									
Pre-BIPA Base Payment X Year FE		X			X			X	
Urban X Year FE			X			X			X
Pre-BIPA Mean of Dep. Var.	30.53	30.53	30.53	484.48	484.48	484.48	485.25	485.25	485.25